



Bayes in the sky: Statistical challenges in cosmology

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(On every Arxiv near you next week)

Bayes in the sky: Bayesian inference and model selection in cosmology

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The application of Bayesian methods in cosmology and astrophysics has flourished over the past decade, spurred by data sets of increasing size and complexity. In many respects, Bayesian methods have proven to be vastly superior to more traditional statistical tools, offering the advantage of higher efficiency and of a consistent conceptual basis for dealing with the problem of induction in the presence of uncertainty. This trend is likely to continue in the future, when the way we collect, manipulate and analyse observations and compare them with theoretical models will assume an even more central role in cosmology.

This review is an introduction to Bayesian methods in cosmology and astrophysics and recent results in the field. I first present Bayesian probability theory and its conceptual underpinnings, Bayes' Theorem and the role of priors. I discuss the problem of parameter inference and its general solution, along with numerical techniques such as Monte Carlo Markov Chain methods. I then review the theory and application of Bayesian model comparison, discussing the notions of Bayesian evidence and effective model complexity, and how to compute and interpret those quantities. Recent developments in cosmological parameter extraction and Bayesian cosmological model building are summarized, highlighting the challenges that lie ahead.

Keywords: Bayesian methods; model comparison; cosmology; parameter inference; data analysis; statistical methods.

The cosmological concordance 'model'



| | | 1D~68% | Best fit | |
|-------------------------------------|-------------------------------------|----------------------------------|----------|--|
| | Cosmological parameters | | | |
| Baryon density | $\Omega_b h^2 \times 10^2$ | $2.23\substack{+0.08\\-0.06}$ | 2.28 | |
| Cold dark matter density | $\Omega_c h^2$ | 0.106 ± 0.004 | 0.107 | |
| Angular size of sound horizon | Θ_* | 1.043 ± 0.003 | 1.042 | |
| Optical depth to reionization | au | $0.084\substack{+0.014\\-0.013}$ | 0.087 | |
| Expansion rate | $H_0 [{\rm Km \ s^{-1} Mpc^{-1}}]$ | 74.3 ± 2.1 | 73.1 | |
| | | | | |
| | Power spectra parameters | | | |
| Amplitude of fluctuations | $\ln(P_s^0\times 10^{10})$ | $3.11\substack{+0.06\\-0.11}$ | 3.15 | |
| Scale dependence of fluctuations | n_0 | 0.973 ± 0.019 | 0.961 | |



- 1. Increasingly complex models and data: "chi-square by eye" simply not enough
- 2. "If it's real, better data will show it": but all the action is in the "discovery zone" around 3-4 sigma significance
- *3. Don't waste time explaining effects which are not there (e.g., reionization at z ~ 16)*
- *4. Plan for the future: which is the best strategy? (survey design & optimization)*
- *5. In some cases, there will be no better data! (cosmic variance)*

Bayesian inference chain



1. Select a model (parameters and range)

- 2. Predict observational signature (as a function of parameters)
- 3. Compare with data

a) derive parameters constraints

PARAMETER INFERENCE

MODEL

COMPARISON

b) compute relative model probability

4. Go back to 1

Challenge #1

Using the right tool for each question

or

How to distinguish between

parameter constraint and model selection tasks

Modelling it all



Primordial fluctuations
A, n_s, dn/dln k, features, ...
10x10 matrix M (isocurvature)
isocurvature tilts, running, ...
Planck scale (B, ω, φ, ...)
Inflation (V, V', V', ...)
Cravity waves (r, n_T, ...)

Astrophysics
Reionization (τ, x_e, history)
Cluster physics
Galaxy formation history

Matter-energy budget
Ω_κ, Ω_Λ, Ω_{cdm}, Ω_{wdm}, Ω_ν, Ω_b
neutrino sector (N_ν, m_ν, c²_{vis}, ...)
dark energy sector (w(z), c_s², ...)
baryons (Y_p, Ω_b)
dark matter sector (b, m_χ, σ, ...)
strings, monopoles, ...

Exotica

Branes, extra dimensions Alignements, Bianchi VII models Quintessence, axions, ...







Parameter inference: (relatively) unproblematic

Prior as "state of knowledge" Different people will have different priors Updated to posterior through the data & Bayes Theorem Will eventually go away as data become better





Likelihood (1 datum)



Posterior after 1 datum



Posterior after 100 data points





Goal: to compare the "performance" of models against the data

$$\mathcal{P}(\boldsymbol{\theta}|\mathbf{d}, \mathcal{M}) = \frac{\mathcal{P}(\mathbf{d}|\boldsymbol{\theta}, \mathcal{M})\pi(\boldsymbol{\theta}|\mathcal{M})}{\mathcal{P}(\mathbf{d}|\mathcal{M})}$$

The model likelihood ("Bayesian evidence")

$$\mathcal{P}(\mathbf{d}|\mathcal{M}) = \int_{\Omega} \mathcal{P}(\mathbf{d}|\boldsymbol{\theta}, \mathcal{M}) \pi(\boldsymbol{\theta}|\mathcal{M}) \mathbf{d}\boldsymbol{\theta}$$

The posterior probability for the model

$$\mathcal{P}(\mathcal{M}|\mathrm{d}) \propto \mathcal{P}(\mathrm{d}|\mathcal{M}) \pi(\mathcal{M})$$

The change in odds is given by the Bayes factor

$$\frac{\mathcal{P}(\mathcal{M}_0|\mathbf{d})}{\mathcal{P}(\mathcal{M}_1|\mathbf{d})} = \frac{\mathcal{P}(\mathbf{d}|\mathcal{M}_0)}{\mathcal{P}(\mathbf{d}|\mathcal{M}_1)} \frac{\pi(\mathcal{M}_0)}{\pi(\mathcal{M}_1)}$$

Jeffreys' scale for the strength of evidence



The Bayes factor
$$B_{01} = \frac{\mathcal{P}(\mathbf{d}|\mathcal{M}_0)}{\mathcal{P}(\mathbf{d}|\mathcal{M}_1)}$$

Interpretation: Jeffreys' scale for the strength of evidence

| In B₀₁ | Odds | Probability (2 models) | #σ | Interpretation |
|--------------------------|---------|---------------------------|--------|-----------------------|
| < 1.0 | < 3:1 | < 0.750 | 1.15 | not worth the mention |
| < 2.5 | < 12:1 | 0.923 | 1.77 | weak |
| < 5.0 | < 150:1 | 0.993 | 2.70 | moderate |
| >5.0 | > 150:1 | > 0.993 | > 2.70 | strong |



Model comparison takes Bayesian inference to the model level - it complements parameter inference

Prior choice is inherent to model specification Gives available model parameter space Related to physical insight into the model

> A model is a choice of parameters and their ranges

$$\mathcal{M} \equiv \{\theta, \pi(\theta|\mathcal{M})\}$$



The Bayes factor balances quality of fit vs extra model complexity. It rewards highly predictive models



- $I = In(prior width / likelihood width) \ge 0$
 - = "wasted" volume of parameter space
 - = amount by which our knowledge has increased

Cosmological applications





Trotta 2007, MNRAS astro-ph/0504022 Trotta 2007, MNRAS Lett, astro-ph/0703063

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ACDM is in the lead

Trotta (2008)

| Competing model | $\Delta N_{\rm par}$ | $\ln B$ | Ref | Data | Outcome |
|---|----------------------------|--|--|---|---|
| Initial conditions Isocurvature modes | | | Bayes factor: In $B < 0$ favours ΛCDM | | |
| CDM isocurvature + arbitrary correlations Neutrino entropy + arbitrary correlations Neutrino velocity + arbitrary correlations | +1 +4 +1 +4 +1 +1 +4 +1 +4 | $\begin{array}{c} -7.6 \\ -1.0 \\ [-2.5, -6.5]^p \\ -1.0 \\ [-2.5, -6.5]^p \\ -1.0 \end{array}$ | [58] [46] [46] [60] [46] | WMAP3+, LSS WMAP1+, LSS, SN Ia WMAP3+, LSS WMAP1+, LSS, SN Ia WMAP3+, LSS WMAP1+, LSS, SN Ia | Strong evidence for adiabaticity Undecided Moderate to strong evidence for adiabaticity Undecided Moderate to strong evidence for adiabaticity Undecided |
| Primordial power spectron No tilt $(n_s = 1)$ | rum —1 | +0.4 $[-1.1, -0.6]^p$ -0.7 -0.9 $[-0.7, -1.7]^{p,d}$ -2.0 -2.6 -2.9 $< -3.9^c$ | [47] [51] [58] [70] [186] [185] [70] [58] [65] | WMAP1+, LSS WMAP1+, LSS WMAP1+, LSS WMAP3+ WMAP3+ WMAP3+, LSS WMAP3+, LSS WMAP3+, LSS | Undecided Undecided Undecided Undecided $n_s = 1$ weakly disfavoured $n_s = 1$ moderately disfavoured $n_s = 1$ moderately disfavoured $n_s = 1$ moderately disfavoured Moderate evidence at best against $n_s \neq 1$ |
| Running Running of running Large scales cut–off | $^{+1}_{+2}_{+2}$ | $[-0.6, 1.0]^{p,d}$ < 0.2^c < 0.4^c $[1.3, 2.2]^{p,d}$ | [186] [166] [166] [186] | WMAP3+, LSS WMAP3+, LSS WMAP3+, LSS WMAP3+, LSS | No evidence for running Running not required Not required Weak support for a cut-off |
| Matter-energy content Non-flat Universe Coupled neutrinos | +1 +1 | -3.8 -3.4 -0.7 | [70] [58] [193] | WMAP3+, HST WMAP3+, LSS, HST WMAP3+, LSS | Flat Universe moderately favoured Flat Universe moderately favoured No evidence for non–SM neutrinos |
| Dark energy sector $w(z) = w_{\text{eff}} \neq -1$ $w(z) = w_0 + w_1 z$ | +1 | $[-1.3, -2.7]^p$ -3.0 -1.1 $[-0.2, -1]^p$ $[-1.6, -2.3]^d$ $[-1.5, -3.4]^p$ -6.0 -1.8 | [187] [50] [51] [188] [187] [50] [187] | SN Ia SN Ia WMAP1+, LSS, SN Ia SN Ia, BAO, WMAP3 SN Ia, GRB SN Ia SN Ia SN Ia SN Ia | Weak to moderate support for Λ Moderate support for Λ Weak support for Λ Undecided Weak support for Λ Weak to moderate support for Λ Strong support for Λ Weak support for Λ |
| $w(z) = w_0 + w_a (1 - a)$ | +2 | $^{-1.8}$ -1.1 $[-1.2, -2.6]^d$ | [188] [189] | SN Ia, BAO, WMAP3 SN Ia, BAO, WMAP3 SN Ia, GRB | Weak support for Λ Weak support for Λ Weak to moderate support for Λ |
| Reionization history No reionization $(\tau = 0)$ No reionization and no tilt | $^{-1}_{-2}$ | $^{-2.6}_{-10.3}$ | [70] [70] | WMAP3+, HST WMAP3+, HST | $\tau \neq 0$ moderately favoured Strongly disfavoured |

Challenge #2

What is a "significant" effect?

or

How not to loose your sleep over 2-sigma "detections" (and why)

Frequentist hypothesis testing



Frequentist hypothesis testing (eg: likelihood ratio) is not what you think it is

A 2-sigma result does not wrongly reject the null hypothesis 5% of the time: at least 29% of 2-sigma results are wrong!

Take an equal mixture of H_0 , H_1 Simulate data, perform hypothesis testing for H_0 Select results rejecting H_0 at 1- α CL What fraction of those results did actually come from H_0 ("true nulls", should not have been rejected)?

| p-value | sigma | fraction of true nulls | lower bound |
|---------|-------|------------------------|-------------|
| 0.05 | 1.96 | 0.51 | 0.29 |
| 0.01 | 2.58 | 0.20 | 0.11 |
| 0.001 | 3.29 | 0.024 | 0.018 |

What went wrong?



For details see: Sellke, Bayarri & Berger, The American Statistician, 55, 1 (2001)

The fundamental mistake is to confuse



Maximising support for the alternative



A Bayesian step is required to obtain the probability of the hypothesis ("model")

When a meaningful prior is not easy to derive, we can still employ an upper bound on the evidence in favour of the new parameter:



Bayesian calibrated p-values



Sellke & Berger (1987), Gordon & Trotta (2007), MNRASLett, arxiv:0706.3014

p-value:
$$\wp = p(t \ge t_{obs}(x)|M_0)$$

For a wide class of unimodal, symmetric priors around ω_0 one can prove that, for all priors

$$B \leqslant \bar{B} = \frac{-1}{\mathrm{e}\wp \ln \wp}$$

where

$$B = B_{10} = \frac{\mathcal{P}(\mathbf{d}|\mathcal{M}_1)}{\mathcal{P}(\mathbf{d}|\mathcal{M}_0)}$$

If the upper bound is small, no other choice of prior will make the extra parameter significant.

A conversion table

Gordon & Trotta (2007), MNRASLett



| Significance p–value | Bayesi evider bounc $ar{B}$ | an Ice I ln $ar{B}$ | sigma | Interpretation (Jeffreys' scale) category |
|-------------------------|--------------------------------------|---------------------------|-------|---|
| 0.05 | 2.5 | 0.9 | 2.0 | |
| 0.04 | 2.9 | 1.0 | 2.1 | 'weak' at best |
| 0.01 | 8.0 | 2.1 | 2.6 | |
| 0.006 | 12 | 2.5 | 2.7 | 'moderate' at best |
| 0.003 | 21 | 3.0 | 3.0 | |
| 0.001 | 53 | 4.0 | 3.3 | |
| 0.0003 | 150 | 5.0 | 3.6 | 'strong' at best |
| 6×10^{-7} | 43000 | 11 | 5.0 | |

Rule of thumb:

a n-sigma result should be interpreted as a n-1 sigma result

Hemispheric asymmetry in the CMB



WMAP1 ILC maps





Eriksen et al (2004)

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 $\Delta \chi^2 = 9$ for 3 extra parameters: is this significant? Bayesian evidence upper bound of 9:1 (weak support)

Challenge #3

Predicting the outcome of future observations

or

How to exploit the known unknowns



Goal: Probability distribution for the outcome of a future observation averaging over current parameters and model uncertainty Multi-model inference: Bayesian generalization of Fisher matrix forecast

- o: current experimente: future experiment
- θ : future max like value

$$P(\theta|o,e) = \sum_{i} P(\theta|o,e,\mathcal{M}^{(i)})P(\mathcal{M}^{(i)}|o)$$

=
$$\sum_{i} P(\mathcal{M}^{(i)}|o) \int P(\theta|\hat{\theta}^{(i)},e,\mathcal{M}^{(i)})P(\hat{\theta}^{(i)}|o,\mathcal{M}^{(i)})d\hat{\theta}^{(i)},$$

current model Fisher matrix present posterior for fiducial posterior (weight)

Application: the PPOD technique Trotta 2007, MNRAS, astro-ph/0703063

Prediction of the value of n_s from Planck data given current data (WMAP3+others): **Predictive Posterior Odds Distribution**



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Prior scaling of PPOD derived analytically from SDDR



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| | | | xford |
|---------------------|--|---|--|
| (Δ_+,Δ) | Required σ for | or evidence level | hysics |
| | moderate | strong | |
| | $(\ln B = 3.0)$ | $(\ln B = 5.0)$ | |
| (0, 10) | 0.4 | $5 \cdot 10^{-2}$ | |
| (2/3, 0) | $3 \cdot 10^{-2}$ | $3\cdot 10^{-3}$ | |
| (0.01, 0.01) | $4 \cdot 10^{-4}$ | $5\cdot 10^{-5}$ | |
| | $(\Delta_+,\Delta) \ (0,10) \ (2/3,0) \ (0.01,0.01)$ | $egin{aligned} & (\Delta_+,\Delta) & 	ext{Required } \sigma \ 	ext{moderate} & \ & (\ln B = 3.0) \ \hline & (0,10) & 0.4 \ & (2/3,0) & 3\cdot 10^{-2} \ & (0.01,0.01) & 4\cdot 10^{-4} \end{aligned}$ | $egin{aligned} & (\Delta_+,\Delta) & 	ext{Required } \sigma 	ext{ for evidence level} \ & 	ext{moderate} & 	ext{strong} \ & (\ln B = 3.0) & (\ln B = 5.0) \ \hline & (0,10) & 0.4 & 5 \cdot 10^{-2} \ & (2/3,0) & 3 \cdot 10^{-2} & 3 \cdot 10^{-3} \ & (0.01,0.01) & 4 \cdot 10^{-4} & 5 \cdot 10^{-5} \ \end{aligned}$ |





Bayesian tools provide a framework for new questions & approaches:

- *Model building:* phenomenologically work which is the "best" model. Needs model insight (prior).
- Experiment design: what is the best strategy to discriminate among models?
- *Performance forecast: how well must we do to reach a certain level of evidence?*
- Science return optimization: use present-day knowledge to optimize future searches